



SOLUTION OF SOME PROBLEMS IN THE THEORY OF PROBABILITIES OF SIGNIFI- CANCE IN AUTOMATIC TELEPHONE EX- CHANGES.

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Summary.—Sections 1–7. First main problem: Systems without waiting arrangements. (Two different presuppositions.) Accompanied by Tables 1, 2, 3, Sections 8–9. Second main problem: Systems with waiting arrangement. (Two different presuppositions.) Accompanied by Tables 4, 5, 6, 7, Sections 10–12. Approximative methods, references, conclusion. Accompanied by Table 8.

(1) *First Main Problem.*—Let us suppose that an automatic system is arranged in such a manner that, somewhere, there are provided x lines to take a certain conversation traffic, *i. e.* the outgoing conversations from a certain number of subscribers. These x lines are said to be co-operative, or to constitute a “group” (or “team”). It is presupposed that all the lines disengaged are accessible. At present we will only speak of systems without waiting arrangements, *i. e.* systems in which the subscriber, when he finds that all x lines are engaged, lays down the receiver, and does not try to get connection again immediately. The probability of thus finding the lines engaged is called the loss, or degree of hindrance, and is here designated by B . With respect to the length of the conversations (sometimes called the holding-time), we will (for the present) suppose

that it is constant, and it will be convenient to consider this quantity equal to 1 ("the natural time-unit"). With respect to the subscribers' calls, it is assumed that they are distributed quite accidentally throughout the time in question (*i. e.* that part of the day when there is most work). This presupposition does not only imply that there may not be points of time within the period of time in consideration at which it may be expected in advance that there will be exceptionally many or few calls, but also that the calls must be mutually independent. In practice these presuppositions will, with great approximation, be fulfilled. The average number of calls per time-unit (intensity of traffic) is called y . The ratio of y to x , *i. e.* the traffic-intensity per line, is designed by a ; it is often called the efficiency of the group. We have to determine B (as function of y and x). The exact expression for this is as follows:

$$B = \frac{\frac{y^x}{x!}}{1 + \frac{y}{1} + \frac{y^2}{2!} + \dots + \frac{y^x}{x!}} \quad . \quad . \quad (1)$$

as shown in the following sections (2-5).

(2) The following proof may be characterised as belonging to the mathematical statistics, and is founded on the conception "statistical equilibrium"—a conception which is of great value in solving certain classes of problems in the theory of probabilities. Let us consider a very great number of simultaneously operating groups of lines of the aforesaid kind (number of lines x , traffic intensity y). If we examine a separate group at a definite moment, we may describe its momentary condition by stating, firstly, how many of the x lines (0, 1, 2, x). are engaged; and secondly, how much there is left of each of the conversations in question. If we examine the same group a short time dt later, we will find that certain changes of two different kinds have taken place. On the one hand, the conversations which were nearly finished will now be over, and the others have become a little older. On the other hand, new calls may have been made, which, however, will have significance only if not all the lines are engaged. (The probability of a new call during the short time dt is ydt .) We suppose that we examine not only one group, but a very great number of groups, both with respect to the momentary condition and the manner in which this alters. The state, of which we thus can get an accurate description, if we use a sufficiently large material, has the characteristic property that, notwithstanding the aforesaid individual alterations, it maintains itself, and, when once begun, remains unaltered, since the alterations of the different kinds balance each other. This property is called "statistic equilibrium."

(3) Temporarily as postulate, we will now set forth the following description of the state of the statistic equilibrium.

The probabilities that 0, 1, 2, 3, x lines are engaged is respectively—

$$\begin{aligned}
 S_0 &= 1 + \frac{y}{1} + \frac{y^2}{2} + \dots + \frac{y^x}{x} \\
 S_1 &= \frac{\frac{y}{1}}{1 + \frac{y}{1} + \frac{y^2}{2} + \dots + \frac{y^x}{x}} \\
 S_2 &= \frac{\frac{y^2}{2}}{1 + \frac{y}{1} + \frac{y^2}{2} + \dots + \frac{y^x}{x}} \quad \dots \quad (2) \\
 &\vdots \\
 S_x &= \frac{\frac{y^x}{x}}{1 + \frac{y}{1} + \frac{y^2}{2} + \dots + \frac{y^x}{x}}
 \end{aligned}$$

where the sum of all the probabilities is 1, as it should be. And we further postulate each of the $x + 1$ aforesaid special conditions, that the still remaining parts of the current conversations (short "remainders") will vary quite accidentally between the limits 0 and 1, so that no special value or combination of values is more probable than the others.

(4) We shall prove that the thus described general state is in statistic equilibrium. For that purpose we must keep account of the fluctuations (increase and decrease), during the time dt , for the $x + 1$ different states, beginning with the first two. The transition from the first state S_0 to the second state S_1 amounts to

$$S_0 \cdot y dt,$$

while the transition from the second S_1 to the first S_0 amounts to

$$S_1 \cdot dt.$$

These quantities are according to (2) equal and thus cancel each other. Furthermore, the amount of transition from S_1 to S_2 is:

$$S_1 \cdot y dt,$$

and, conversely, the transition from S_2 to S_1 is:

$$S_2 \cdot 2 \cdot dt,$$

which two quantities also cancel each other.

Finally, we have

$$S_{x-1} \cdot y \cdot dt$$

and

$$S_x \cdot x \cdot dt,$$

which also cancel each other. The result is that the reciprocal changes which take place between the $x + 1$ different conditions during the time dt , compensate each other, so that the distribution remains unaltered. We still have to prove that neither will there be any alterations in the distribution of the magnitude of the remainders, *i. e.* that the decrease and increase, also in this respect, compensate each other.

(5) Let us consider the cases in which the number of current conversations is n , and among these cases, more especially those in which the magnitudes of the n remainders lie, respectively, between the following limits :

$$\begin{array}{ccccccc} t_1 & \text{and} & t_1 + \Delta_1, \\ t_2 & \text{and} & t_2 + \Delta_2, \\ : & : & : & : & : & : \\ t_n & \text{and} & t_n + \Delta_n. \end{array}$$

The probability for this is (according to Section 3) :

$$\Delta_1 \cdot \Delta_2 \cdot \Delta_3 \dots \Delta_n \cdot S_n.$$

During the time dt there may occur both increase and decrease, and this may occur in four different ways.

Firstly, transition to S_{n+1} ; namely, if there comes a call; the probability for this will be :

$$\Delta_1 \cdot \Delta_2 \cdot \Delta_3 \dots \Delta_n \cdot S_n \cdot y \cdot dt.$$

Secondly, transition from S_{n+1} ; namely, if one among the $n + 1$ current conversations finishes during the time dt , and, thereafter, the n remainders lie between the above settled limits. The corresponding probability is :

$$\Delta_1 \cdot \Delta_2 \cdot \Delta_3 \dots \Delta_n (n+1) S_{n+1} \cdot dt,$$

which is equal to the preceding one.

Thirdly, transition from S_n itself; namely, if, among the n remainders, the $n - 1$ lie between the settled limits, and the one lies just below the lower limit in question, at a distance shorter than dt . The probability for this will be :

$$\Delta_1 \cdot \Delta_2 \cdot \Delta_3 \dots \Delta_n \left(\frac{1}{\Delta_1} + \frac{1}{\Delta_2} + \dots + \frac{1}{\Delta_n} \right) S_n \cdot dt.$$

Fourthly, transition to S_n itself; namely, if, among the n remainders, the $n - 1$ lie between the settled limits, and the one lies just below the upper limit, at a distance shorter than dt . The probability of this eventuality is obviously equal to the preceding one. Thus, there is a balance. So it is proved by this that there will be statistical equilibrium. On the other hand, every other

supposition than the one set forth in Section 3 will at once be seen to be inconsistent with statistic equilibrium. The formulæ in Section 3 are now proved, and thereby the proposition in Section 1 is also proved.

(6) The above presupposition, that all conversations are of equal length, applies with great approximation to trunk-line conversations, but, of course, not to the usual local conversations. Now, a statistic investigation, which I have undertaken in Copenhagen, has shown that the length of these conversations very nearly obeys a simple law of distribution, which may be expressed as follows:

The probability that the length will exceed n is equal to

$$e^{-n},$$

when the average length is taken to be equal to 1. Or, in other words, the probability that a conversation which has continued for some time is nearly finished, is quite independent of the length of the time which has already elapsed. The average number of finishing conversations during the time dt per current conversation becomes always equal to dt . It is easy now to see, that we must arrive at the same expression (1) for B as under the former presupposition, only that the proof becomes somewhat simpler, because it is only necessary to take into account the number of current conversations without paying any attention to their age. (The two aforesaid presuppositions do not lead to the same result in all problems.)

(7) In the Table 1 are shown the numerical values of the "loss" B as dependent of x and y (or u), and as given by the proposed theory.

In the Table 2 the results of formula (1) are presented in another form, which is probably the one that is most useful in practice; x and B are here entry numbers, and the table gives y as function of x and B .

In Table 3A only the first and second line treat of systems with "pure" groups (to which formula No. 1 applies). The values given in the third line correspond to a different system, in which a special arrangement the so-called "grating and inter-connecting" is used. We may describe this arrangement as follows:

The number of contacts of the selectors (here ten) is less than the number of lines (here eighteen) in the "group." Thus each call searches not all eighteen but only ten lines. It is hereby presupposed (for the sake of simplicity) that the ten lines are each time accidentally chosen, out of the eighteen, and that they are tested one after the other in an accidental order. The method of calculation here to be used may be considered as a natural extension of the method which leads to formula (1), but it is, of course, a little more

complicated. A few results of this kind of calculations may be seen in the two Tables 3A and 3B. Finally, it may be remarked that the systems for "grating and interconnecting" presently in practical use (namely, the systems of Siemens and Halske), which I, however, do not know in detail, are said to deviate a little from the description given here, and, therefore, it may be expected that they will give somewhat less favourable results.

(8) *Second Main Problem.*—The problem which we are now to treat concerns systems with waiting arrangements. Here the problem to be solved is determining the probability $S(<n)$ of a waiting time greater than an arbitrary number n , greater than or equal to zero. The last case is the one which is most frequently asked for. In the same manner we define $S(<n)$ where $S(<n) + S(>n) = 1$. Furthermore, we may ask about the average waiting time M . We shall answer these questions in the following. We may begin, even here, by assuming that the length of the conversations is constant and equal to 1. The accurate treatment of this case gives rise to quite difficult calculations, which, however, are unavoidable in this case. Among other things, we find that we cannot use the same formula for $S(<n)$ for all values of n , but we must distinguish between the various successive "periods," or spaces of time of the length 1. In practice, however, the first period will, as a rule, be the most important. I shall content myself by giving, without proof, the necessary formulæ for the cases $x = 1, 2$, and 3, and then (chiefly for the purpose of showing the possibility of carrying out the practical calculation) the corresponding numerical results. Formulæ and results for $x = 1$ I have already given in an article in 'Nyt Tidsskrift for Matematik,' B. 20, 1909. The formulæ for greater values of x , i. e. $x = 10, x = 20$ are quite analogous to those here given. The numerical tables corresponding hereto are in course of calculation, but are not yet finished.

(9) There still remains the problem of investigating the magnitude of the waiting times in systems with waiting arrangement under the second presupposition, namely, that the lengths of the talks vary in the manner already described in Section 6.

Here we find, without difficulty, the following two formulæ:

$$S(>0) = c. \quad (3)$$

$$S(>n) = c \cdot e^{-x-yin} \quad (4)$$

where

$$C = \frac{1}{1 + \frac{y}{1} + \frac{y^2}{2} + \dots + \frac{y^{x-1}}{x-1} + \frac{y^x}{x} \cdot \frac{x}{x-y}} \quad (5)$$

while x and y have the same significance as before, and the average

length of a conversation is equal to 1. The second formula gives the correct values for $n \geq 0$.

For the average waiting time we get the formula :

$$M = \frac{c}{x-y} \quad (6)$$

The numerical calculation causes no special difficulty. It ought, perhaps, to be remarked that, both here and in Section 8, it is supposed that the waiting calls are despatched in the order in which they have been received. If this does not take place in practice, it will, of course, as a rule, have a slight effect upon the value of S ($>n$), but not at all on the value of M , neither on S (>0).

(10) *Approximate Formulæ.*—The *exact* formulæ given above are throughout so convenient, that there is hardly any need of approximate formulæ. This does not, however, apply to the formulæ which concern the second main problem, first presupposition. Therefore, it may be worth while to mention a couple of approximate methods which quickly lead to a serviceable result, at least in such cases as have the greatest practical significance.

One of these methods has already been used by me, by request of Mr. P. V. Christensen, Assistant Chief-Engineer of the Copenhagen Telephone Company, for calculating the tables given in the first pages of his fundamental work, "The Number of Selectors in Automatic Telephone Systems" (published in the POST OFFICE ELECTRICAL ENGINEERS' JOURNAL, October, 1914; also in 'Elektroteknikeren,' No. 48, 1913; 'E. T. Z. H.,' 46, 1913).

The second of the tables given by Mr. Christensen in the named place is reproduced, in a somewhat altered arrangement, as Table 8. Since the method used has hardly been given in full, I shall here say a few words about the same. The probability of just x calls during a period of time for which the average number is y , is, as well known, under usual presuppositions (Section 1) :

$$e^{-y} \frac{y^x}{x!}$$

The mathematical proposition here used is due to S. D. Poisson ('Recherches sur la probabilité, etc.,' 1837), and has later been studied by L. v. Bortkewitsch ('Das Gesetz der kleinen Zahlen,' 1898). The function is tabulated by the latter (*loc. cit.*), and later by H. E. Soper ('Biometrika,' vol. x, 1914; also in K. Pearson ('Tables for Statisticians, etc.,' 1914).

Thus the probability of x or more calls during the mentioned period of time is :

$$P = e^{-y} \frac{y^x}{x!} + e^{-y} \frac{y^{x+1}}{(x+1)!} + e^{-y} \frac{y^{x+2}}{(x+2)!} + \dots \quad (7)$$

For our present purpose we must consider such a period of time, lying immediately before the special call which we just are interested in.

It will then be seen that P , in many cases, namely, when y is not unproportionally great, will be a good approximate value for the fraction of the calls which will find all the lines engaged (or "the probability of not getting through"). Thus P in the case of exchanges without waiting arrangements approximates the "loss," and here gives obviously a somewhat too great value. In exchanges with waiting arrangement P approximates the quantity $S (>0)$, the probability of delay, and gives here a somewhat too small value. Or, if it is the fraction named above which is given in beforehand, as is generally the case in practice, where often the value 0.001 is used, the formula will show the connection between y and x . The values of y found in this manner (see Table 8) will never deviate 5 per cent. from the correct values in systems without waiting arrangement; never 1 per cent. in systems with waiting arrangement (both presuppositions), if we take, as an instance, the named frequently used value $P = 0.001$. Intermediate systems between the two main classes of exchanges may, of course, be treated with good result according to the same method.

If, in systems with waiting arrangement, we ask about the number of waiting times beyond a certain limit n , $S (>n)$, an extension of the same formula may be used, y being replaced by $y(1-n)$. The method is best suited to small values of n , and the error goes to the same side as above mentioned. Furthermore, it may be mentioned in this connection that if we use, in the case considered, the formula following from presupposition No. 2, instead of those based upon presupposition No. 1, as a rule, the errors thus introduced will be small; they go, this time, in such a direction that we get too great values for $S (>0)$ and $S (<n)$; or, if it is y which is sought, there will be too small values for y .

(11) It will be too lengthy to describe or mention, in this connection, all the systematic practical experiments and measurements (also only partly published), which of late years have been made, partly by the firms in question (namely, Siemens and Halske, and Western Electric Co.), partly by others, or such purely empirical formulæ which have thus been set forth. On the other hand, it would be incorrect to neglect one or two interesting theoretical works from latter years, which directly concern one of the above-treated problems. In his doctor's thesis, Mr. F. Spiecker ('Die Abhängigkeit des erfolgreichen Fernsprechanrufes von der Anzahl der Verbindungsorgane,' 1913), has given a method for determining the loss in systems without waiting arrangement, which (as he himself remarks) is not quite free from errors, and which, besides,

is so complicated that it can hardly be taken into consideration in practical work. It ought to be noticed, however, that the results in the cases in which the author has carried through his calculations lie very near the results of formula No. 1 given above. In the same work there is also given an approximative formula, which can best be compared with the formula for P (Section 10 above). The difference is due only to a somewhat deviating, and probably less practical, formulation of the problem. Mr. W. H. Grinsted, in his treatise, "A Study of Telephone Traffic Problems, etc." (POST OFFICE ELECTRICAL ENGINEERS' JOURNAL, April, 1915), presents a solution of the same problem. Since this solution has, if I do not err, by many readers as well as by the author himself, been considered mathematically correct, it may be here remarked that in the derivation of the formula in question there has occurred an error, and that, for this reason, the formula gives rather incorrect results. It should be added that this treatise is a reprint of an older work from 1907 (which I have not been able to examine). In spite of the faulty results, the work of Grinsted is, however, when its time of publication is considered, of no little merit.

(12) In closing this article, I feel called upon to render my best thanks to Mr. F. Johannsen, Managing Director of the Copenhagen Telephone Co., not only for his interest in the investigations here recorded, but also for his energetic initiative in starting rational and scientific treatment of many different problems in connection with telephone traffic. I also owe many thanks to Mr. J. L. W. V. Jensen, Engineer-in-Chief to the same Company, for his valuable assistance especially in the treatment of some mathematical difficulties.

[Owing to lack of space we are unable to publish the various tables quoted in the foregoing. Should any of our readers desire to pursue the study, we shall be pleased to supply a copy of them on loan. Mr. Grinsted has seen the above (11), and hopes to comment upon it in a future issue after he has received details of Mr. Erlang's criticism of his formulæ—EDS. POST OFFICE ELECTRICAL ENGINEERS' JOURNAL.]